

02/26/04 Thurs MC is irreducible, aperiodic,
finite state

$$\therefore \lim_{n \rightarrow \infty} \text{Prob}(X_n = j) = \pi_j$$

Where $\vec{\pi}$ is the unique stat dist.

Let's calculate $\vec{\pi}$ then use for questions.

$$\vec{\pi} \cdot P = \vec{\pi} \quad \text{with} \quad \sum_{j=0}^M \pi_j = 1$$

~~0th eqn:~~

$$0\text{th eqn: } p \sum_{j=0}^M \pi_j = \pi_0$$

$$1 \leq j \leq M-1: (1-p) \pi_{j-1} = \pi_j$$

$$M\text{th eqn: } (1-p)(\pi_{M-1} + \pi_M) = \pi_M$$

$$\text{Middle eqns: } \pi_j = (1-p)^j \pi_0 \quad \text{for } 1 \leq j \leq M-1$$

$$0\text{th eqn: } \pi_0 = p$$

$$\pi_j = p(1-p)^j$$

$$\pi_M = 1 - \sum_{j=0}^{M-1} \pi_j = 1 - p \sum_{j=0}^{M-1} (1-p)^j$$

$$= 1 - p \left(\frac{1 - (1-p)^M}{p} \right) = (1-p)^M$$

This describes $\vec{\pi}$.

Consider the question of how many products are inspected first.

I = fraction of items inspected

$$= \lim_{n \rightarrow \infty} \frac{\# \text{ inspections } (n)}{\# \text{ products that pass through in } n \text{ inspections.}}$$

T_n = # products passing the inspection station between inspection # n and inspection # $n+1$

$$T_n = f(X_n) \begin{cases} 1 & \text{for } X_n \neq M \\ r & \text{for } X_n = M \end{cases}$$

$$I = \lim_{n \rightarrow \infty} \frac{n}{\sum_{j=1}^n T_j}$$

Probability theory: Law of Large Numbers

If $\{Y_j\}_{j=0}^{\infty}$ are iid, n r.v. variables

then $\frac{1}{n} \sum_{j=0}^n Y_j \xrightarrow{n \rightarrow \infty} \langle Y \rangle$ almost surely
(prob = 1)

Law of Large Numbers for MC (Resnick 2.1a)

If $\{Y_n\}_{n=0}^{\infty}$ is a F.S.H.D.T MC

which is ergodic (aperiodic + irreducible + ~~recurrent~~ (positive recurrent))

then for deterministic functions f

$$\frac{1}{N} \sum_{n=1}^N f(Y_n) \rightarrow \sum_{j \in S} \pi_j^{(Y)} f(j)$$

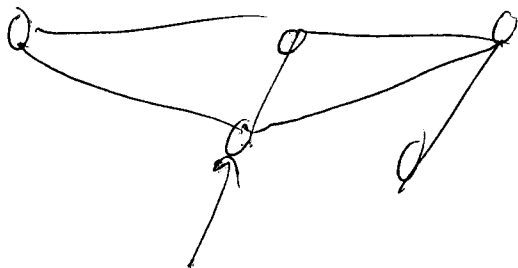
almost surely (w/ prob. 1)

Where π is stat. dist.

Idea of proof: Reduce to LLN for iid rvs.

- dissection principle (Resnick):

The trajectories of MC between successive visits to a reference state are independent of each other



Apply the LLN for MC

$$I = \lim_{h \rightarrow \infty} \frac{n}{\sum_{j=1}^n T_j} = \frac{1}{\lim_{n \rightarrow \infty} \frac{\sum_{j=1}^n T_j}{n}}$$

$$= \frac{1}{\sum_{j=0}^M \pi_j f(j)}$$

$$f(j) = \begin{cases} 1 & \text{for } 0 \leq j \leq M-1 \\ r & \text{for } j = M \end{cases}$$

$$= \frac{1}{\left(\sum_{j=0}^{M-1} \pi_j \right) 1 + \pi_M r} = \frac{1}{1 - \pi_M + \pi_M r}$$

$$= \frac{1}{1 + (r-1)\pi_M} \quad \text{~~is not correct~~$$

$$I = \frac{1}{1 + (r-1)(1-p)^M} = \text{fraction of products inspected}$$

How about fraction of shipped products that are defective?

$$D = \frac{\# \text{ products shipped + defective}}{\# \text{ products shipped}}$$

$$= \frac{\sum_{n=1}^N D_n}{\sum_{n=1}^N S_n}$$

$D_n = \#$ defective products shipped + defective right from ~~from~~ ~~the~~ n th inspection up to $(n+1)$ st inspection (but not including)

$$D_n = 0 \quad \text{when} \quad 0 \leq X_n \leq M-1$$

D_n is random when $X_n = M$!

Can't write D_n as a deterministic function of X_n .

$$\text{Prob}(D_n = d) = \begin{cases} \delta_{d0} & \text{when } X_n \neq M \\ b(d; r-1, p) & \text{when } X_n = M \end{cases}$$

$$\delta_{ij} = \begin{cases} 1 & \text{for } i=j \\ 0 & \text{for } i \neq j \end{cases}$$

binomial dist.

$S_n = \#$ products shipped ~~to~~ from
 n th inspection up to, but not
including inspection $\#$ $n+1$

$$S_n = 0 \quad \text{if} \quad \cancel{X_n = 0} \quad X_n = 0$$

$$S_n = 1 \quad \text{if} \quad 1 \leq X_n \leq M-1$$

$$S_n = r \quad \text{if} \quad X_n = M$$

$$S_n = g(X_n)$$

$$g(j) = \begin{cases} 0 & \text{for } j = 0 \\ 1 & \text{for } 1 \leq j \leq M-1 \\ r & \text{for } j = M \end{cases}$$

$$D = \lim_{N \rightarrow \infty} \frac{\sum_{n=1}^N D_n}{\sum_{n=1}^N S_n} = \lim_{N \rightarrow \infty} \frac{\frac{1}{N} \sum_{n=1}^N D_n}{\frac{1}{N} \sum_{n=1}^N S_n}$$

LLN for MC:

$$\frac{1}{N} \sum_{n=1}^N S_n \xrightarrow{N \rightarrow \infty} \sum_{j=0}^M \pi_j g(j) \quad \text{w/ probs. } \pi_j$$

By viewing the extended MC $\{X_n, D_n\}$
and apply LLN

$$\frac{1}{N} \sum_{n=1}^N D_n \xrightarrow{N \rightarrow \infty} \sum_{j=0}^M \pi_j \langle D_n | X_n = j \rangle$$

$$\langle D_n | X_n = j \rangle = 0 \quad \text{for } 0 \leq j \leq M-1$$

$$\langle D_n | X_n = M \rangle = (r-1)p \quad \text{for } M$$

$$D = \frac{\sum_{j=0}^M \pi_j \langle D_n | X_n = j \rangle}{\sum_{j=0}^M \pi_j g(j)}$$

$$= \frac{\left(\sum_{j=0}^{M-1} \pi_j \right) 0 + \pi_M (r-1)p}{\pi_0 (0) + \left(\sum_{j=1}^M \pi_j \right) 1 + \pi_M r}$$

$$= \frac{\pi_M (r-1)p}{1 - \pi_M - \pi_M r} = \frac{\pi_M (r-1)p}{1 + (r-1)\pi_M - p}$$

$$= \frac{(1-p)^M (r-1)p}{1-p + (r-1)(1-p)^M} = \frac{(r-1)p (1-p)^{M+1}}{1 + (r-1)(1-p)^{M-1}}$$

How is D related to I ?

$$(r-1)(1-p)^M = \frac{1-I}{I} \quad \text{by inserting the eqn for } I$$

Plugging into D

$$D = \frac{p \frac{1-I}{I}}{1-p + \frac{1-I}{I}} = \frac{p(1-I)}{1-pI}$$

Independent of $M+r$!

→ Relates outgoing quality to effort,

How does this compare w/ ~~with~~ simple random sampling where each product is inspected w/ prob. I ?

Fraction of products shipped

$$= \text{Prob}(\text{product is inspected + good, or not inspected})$$

$$= I(1-p) + 1-I = 1-Ip$$

Fraction of products that are defective + shipped

$$= \text{Prob}(\text{product not inspected and bad})$$

$$= (1-I)p$$

For simple random sampling,

$$D = \frac{(1-I)p}{1-Ip} = \frac{p-pI}{1-Ip}$$

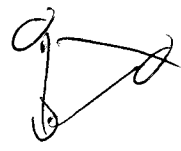
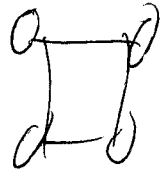
Same as for regular adaptive sampling scheme!

No benefit because products defects independent.

- may do better for clumped defects \rightarrow Homework!

What about FS # DT MC which are not irreducible?

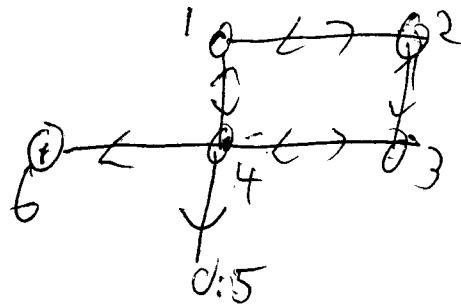
Trivial: Disconnected classes



break into separate MC's.

More interesting is MC w/ 1-way paths between communication classes.

Random walk w/ absorbing BC



$\{1, 2, 3, 4\}$: Transient

$\{5\}$: Recurrent

$\{6\}$: Recurrent

For reducible MC's: canonical decomposition

Classification of states:

A state $j \in S$ is recurrent if

the MC starting from state j has prob = 1 to return to state j .

A state $j \in S$ is transient if

the MC starting from state j

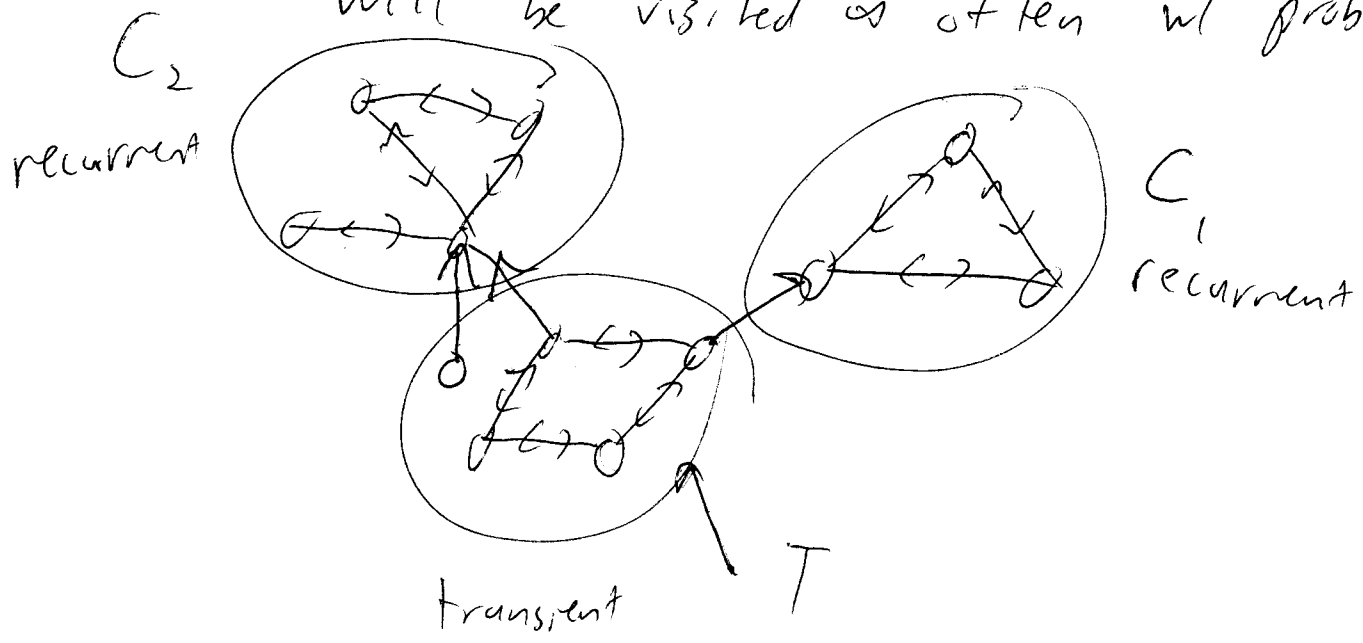
has prob < 1 to return to state j .

Facts: 1) Recurrence + transience are class properties (same for all states in communication class)

2) In a recurrent class C_k then state

2) Recurrent state will be visited as often (w/ prob. 1) if MC starts in ^{that} state.

3) In a recurrent class C_k then starting from any state in that class, any other state will be visited as often w/ prob. 1.



Decompose the MC into

$$S = T \cup_k C_k$$

↑
all transient

states (even if not 1 class)

Each C_k is
a recurrent
class,

If you start in a recurrent class,
long-time properties can be calculated
by just treating the recurrent class
as its own irreducible MC